# Bayesian Nonparametrics 

Part I

Peter Orbanz

## Overview

## Today

1. Basic terminology
2. Clustering
3. Latent feature models

## Tomorrow

5. Constructing nonparametric Bayesian models
6. Exchangeability
7. Asymptotics

## Parameters and Patterns

Parameters

$$
P(X \mid \theta) \quad=\quad \text { Probability }[\text { data } \mid \text { pattern }]
$$




Inference idea

$$
\text { data }=\text { underlying pattern }+ \text { independent noise }
$$

## TERMINOLOGY

## Parametric model

- Number of parameters fixed (or constantly bounded) w.r.t. sample size


## Nonparametric model

- Number of parameters grows with sample size
- $\infty$-dimensional parameter space

Example: Density estimation


Parametric


Nonparametric

## NONPARAMETRIC BAYESIAN MODEL

## Definition

A nonparametric Bayesian model is a Bayesian model on an $\infty$-dimensional parameter space.

## Interpretation

Parameter space $\mathcal{T}=$ set of possible patterns, for example:

| Problem | $\mathcal{T}$ |
| :---: | :---: |
| Density estimation | Probability distributions |
| Regression | Smooth functions |
| Clustering | Partitions |

Solution to Bayesian problem $=$ posterior distribution on patterns

## EXCHANGEABILITY

## Can we justify our assumptions?

Recall:

$$
\text { data }=\text { pattern }+ \text { noise }
$$

In Bayes' theorem:

$$
Q\left(d \theta \mid x_{1}, \ldots, x_{n}\right)=\frac{\prod_{j=1}^{n} p\left(x_{j} \mid \theta\right)}{p\left(x_{1}, \ldots, x_{n}\right)} Q(d \theta)
$$



## Definition

$X_{1}, X_{2}, \ldots$ are exchangeable if $P\left(X_{1}, X_{2}, \ldots\right)$ is invariant under any permutation $\sigma$ :

$$
P\left(X_{1}=x_{1}, X_{2}=x_{2}, \ldots\right)=P\left(X_{1}=x_{\sigma(1)}, X_{2}=x_{\sigma(2)}, \ldots\right)
$$

In words:
Order of observations does not matter.

## Exchangeability and Conditional Independence

## De Finetti's Theorem

$$
\begin{gathered}
P\left(X_{1}=x_{1}, X_{2}=x_{2}, \ldots\right)=\int_{\mathbf{M}(\mathcal{X})}\left(\prod_{j=1}^{\infty} \theta\left(X_{j}=x_{j}\right)\right) Q(d \theta) \\
\Uparrow \\
X_{1}, X_{2}, \ldots \text { exchangeable }
\end{gathered}
$$

where:

- $\mathbf{M}(\mathcal{X})$ is the set of probability measures on $\mathcal{X}$
- $\theta$ are values of a random probability measure $\Theta$ with distribution $Q$


## Implications

- Exchangeable data decomposes into pattern and noise
- More general than i.i.d.-assumption
- Caution: $\theta$ is in general an $\infty$-dimensional quantity


## Clustering

## Clustering

- Observations $X_{1}, X_{2}, \ldots$
- Each observation belongs to exactly one cluster
- Unknown pattern $=$ partition of $\{1, \ldots, n\}$ or $\mathbb{N}$



## Mixture models

Mixture models

$$
p(x \mid m)=\int_{\Omega_{\theta}} p(x \mid \theta) m(d \theta)
$$

$m$ is called the mixing measure
Two-stage sampling
Sample $X \sim p(. \mid m)$ as:

1. $\Theta \sim m$
2. $X \sim p(. \mid \theta)$

Finite mixture model

$$
p(x \mid \boldsymbol{\theta}, \mathbf{c})=\int_{\Omega_{\theta}} p(x \mid \theta) m(d \theta) \quad \text { with } \quad m(.)=\sum_{k=1}^{K} c_{k} \delta_{\theta_{k}}(.)
$$

## BAYESIAN MM

Random mixing measure

$$
M(.)=\sum_{k=1}^{K} C_{k} \delta_{\Theta_{k}}(.)
$$

Conjugate priors
A Bayesian model is conjugate if the posterior is an element of the same class of distributions as the prior ("closure under sampling").

| $p(x \mid \theta)$ | conjugate prior |
| :---: | :---: |
| $\frac{1}{\frac{1}{Z(\theta)} h(x) \exp (\langle S(x), \theta\rangle)}$ | $\frac{1}{K(\lambda, y)} \exp (\langle\theta, y\rangle-\lambda \log Z(\theta))$ |
| Gaussian | Gaussian/inverse Wishart |
| multinomial | Dirichlet |
| $\ldots$ | $\ldots$ |

- Choose conjugate prior for each parameter
- In particular: Dirichlet prior on $\left(C_{1}, \ldots, C_{k}\right)$


## DIRICHLET PROCESS MIXTURES

## Dirichlet process

A Dirichlet process is a distribution on random probability measures of the form

$$
M(.)=\sum_{k=1}^{\infty} C_{k} \delta_{\Theta_{k}}(.) \quad \text { where } \quad \sum_{k=1}^{\infty} C_{k}=1
$$

## Constructive definition of $\operatorname{DP}\left(\alpha, G_{0}\right)$

$$
\begin{aligned}
\Theta_{k} & \sim_{\mathrm{idd}} G_{0} \\
V_{k} & \sim_{\mathrm{idd}} \operatorname{Beta}(1, \alpha)
\end{aligned}
$$

Compute $C_{k}$ as

$$
C_{k}:=V_{k} \prod_{i=1}^{k-1}\left(1-V_{i}\right)
$$

"Stick-breaking construction"

## Posterior distribution

DP Posterior

$$
\theta_{n+1} \mid \theta_{1}, \ldots, \theta_{n} \sim \frac{1}{n+\alpha} \sum_{j=1}^{n} \delta_{\theta_{j}}\left(\theta_{n+1}\right)+\frac{\alpha}{n+\alpha} G_{0}\left(\theta_{n+1}\right)
$$

## Mixture Posterior

$$
p\left(x_{n+1} \mid x_{1}, \ldots, x_{n}\right)=\sum_{k=1}^{K_{n}} \frac{n_{k}}{n+\alpha} p\left(x_{n+1} \mid \theta_{k}^{*}\right)+\frac{\alpha}{n+\alpha} \int p\left(x_{n+1} \mid \theta\right) G_{0}(\theta) d \theta
$$

Conjugacy

- The posterior of DP $\left(\alpha, G_{0}\right)$ is $\operatorname{DP}\left(\alpha+n, \frac{1}{n+\alpha}\left(\sum_{k} n_{k} \delta_{\theta_{k}^{*}}+\alpha G_{0}\right)\right)$
- Hence: The Dirichlet process is conjugate.


## Inference

## Latent variables

$$
p\left(x_{n+1} \mid x_{1}, \ldots, x_{n}\right)=\sum_{k=1}^{K_{n}} \frac{n_{k}}{n+\alpha} p\left(x_{n+1} \mid \theta_{k}^{*}\right)+\frac{\alpha}{n+\alpha} \int p\left(x_{n+1} \mid \theta\right) G_{0}(\theta) d \theta
$$

We do not actually observe the $\Theta_{j}$ (they are latent). We observe $X_{j}$.

## Assignment probabilities

$$
\left(\begin{array}{cccc}
q_{10} & q_{11} & \ldots & q_{1 K_{n}} \\
\vdots & \vdots & & \vdots \\
q_{n 0} & q_{n 1} & \ldots & q_{n K_{n}}
\end{array}\right)
$$

Where:

- $q_{j k} \propto n_{k} p\left(x_{j} \mid \theta_{k}^{*}\right)$
- $q_{j 0} \propto \alpha \int p\left(x_{j} \mid \theta\right) G_{0}(\theta) d \theta$


## Gibbs Sampling

Uses an assignment variable $\phi_{j}$ for each observation $X_{j}$.

- Assignment step: Sample $\phi_{j} \sim \operatorname{Multinomial}\left(q_{j 0}, \ldots, q_{j K_{n}}\right)$
- Parameter sampling: $\theta_{k}^{*} \sim G_{0}\left(\theta_{k}^{*}\right) \prod_{x_{j} \in \operatorname{Cluster} k} p\left(x_{j} \mid \theta_{k}^{*}\right)$


## Number of Clusters

## Dirichlet process

$K_{n}=\#$ clusters in sample of size $n$

$$
\mathbb{E}\left[K_{n}\right]=O(\log (n))
$$

## Modeling assumption



- Parametric clustering: $K_{\infty}$ is finite (possibly unknown, but fixed).
- Nonparametric clustering: $K_{\infty}$ is infinite


## Rephrasing the question

- Estimate of $K_{n}$ is controlled by distribution of the cluster sizes $C_{k}$ in $\sum_{k} C_{k} \delta_{\Theta_{k}}$.
- Ask instead: What should we assume about the distribution of $C_{k}$ ?


## GENERALIZING THE DP

## Pitman-Yor process

$$
p\left(x_{n+1} \mid x_{1}, \ldots, x_{n}\right)=\sum_{k=1}^{K_{n}} \frac{n_{k}-d}{n+\alpha} p\left(x_{n+1} \mid \theta_{k}^{*}\right)+\frac{\alpha+K_{n} \cdot d}{n+\alpha} \int p\left(x_{n+1} \mid \theta\right) G_{0}(\theta) d \theta
$$

Discount parameter $d \in[0,1]$.
Cluster sizes



## Power Laws

The distribution of cluster sizes is called a power law if

$$
C_{j} \sim \gamma(\beta) \cdot j^{-\beta} \quad \text { for some } \beta \in[0,1]
$$

Examples of power laws

- Word frequencies
- Popularity (number of friends) in social networks

Pitman-Yor language model


## RANDOM PARTITIONS

## Discrete measures and partitions

Sampling from a discrete measure determines a partition of $\mathbb{N}$ into blocks $b_{k}$ :

$$
\Theta_{n} \sim_{\mathrm{iid}} \sum_{k=1}^{\infty} c_{k} \delta_{\theta_{k}^{*}} \quad \text { and set } \quad n \in b_{k} \quad \Leftrightarrow \quad \Theta_{n}=\theta_{k}^{*}
$$

As $n \longrightarrow \infty$, the block proportions converge: $\frac{\left|b_{k}\right|}{n} \longrightarrow c_{k}$

## Induced random partition

The distribution of a random discrete measure $M=\sum_{k=1}^{\infty} C_{k} \delta_{\Theta_{k}}$ induces the distribution of a random partition $\Pi=\left(B_{1}, B_{2}, \ldots\right)$.

## Exchangeable random partitions

- $\Pi$ is called exchangeable if its distribution depends only on the sizes of its blocks.
- All exchangeable random parititions, and only those, can be represented by a random discrete distribution as above (Kingman's theorem).


## Chinese Restaurant Process

## Chinese Restaurant Process

The distribution of the random partition induced by the Dirichlet process is called the Chinese Restaurant Process.
"Customers and tables" analogy


Customers $=$ observations $($ indices in $\mathbb{N})$
Tables $=$ clusters (blocks)

## Historical remark

- Originally introduced by Dubins \& Pitman as a distribution on infinite permutations
- A permutation of $n$ items defines a partition of $\{1, \ldots, n\}$ (regard cycles of permutation as blocks of partition)
- The induced distribution on partitions is the CRP we use in clustering


## Families of Exchangeable Random Partitions



## Random Discrete Measures

## Classification (due to Prünster)

| class | probability of new cluster | prior class |
| :---: | :--- | :--- |
| I | $\mathbb{P}\left\{\Theta_{n+1} \in\right.$ new cluster $\left.\mid \Theta^{(n)}\right\}=f(n)$ | Dirichlet processes |
| II | $\mathbb{P}\left\{\Theta_{n+1} \in\right.$ new cluster $\left.\mid \Theta^{(n)}\right\}=f\left(n, K_{n}\right)$ | Gibbs-type measures |
| III | $\mathbb{P}\left\{\Theta_{n+1} \in\right.$ new cluster $\left.\mid \Theta^{(n)}\right\}=f\left(n, K_{n}, \mathbf{n}\right)$ |  |

## General partition priors

- Gibbs-type measures are completely classified [GP06b]
- Properties of some cases well-studied, e.g.:
- Dirichlet process
- Pitman-Yor process
- Normalized inverse Gaussian process [LMP05b]
- In the future: We will have a range of models which express different prior assumptions on the distribution of cluster sizes.


## Summary: Clustering

## Nonparametric Bayesian clustering

- Infinite number of clusters, $K_{n} \leq n$ of which are observed.
- If partition exchangeable, it can be represented by a random discrete distribution.


## Inference

Latent variable algorithms, since assignments ( $\equiv$ partition) not observed.

- Gibbs sampling
- Variational algorithms


## Prior assumption

- Distribution of cluster sizes.
- Implies prior assumption on number $K_{n}$ of clusters.


## Latent Feature Models

## Indian Buffet process

## Latent feature models

- Grouping problem with overlapping clusters.
- Encode as binary matrix: Observation $n$ in cluster $k \quad \Leftrightarrow \quad X_{n k}=1$
- Alternatively: Item $n$ possesses feature $k \quad \Leftrightarrow \quad X_{n k}=1$


## Indian buffet process (IBP)

1. Customer 1 tries Poisson $(\alpha)$ dishes.
2. Subsequent customer $n+1$ :

- tries a previously tried dish $k$ with probability $\frac{n_{k}}{n+1}$,
- tries Poisson $\left(\frac{\alpha}{n+1}\right)$ new dishes.


## Properties

- An exchangeable distribution over finite sets (of dishes).
- Intepretation:

Observation (= customer) $n$ in cluster (= dish) $k$ if customer "tries dish $k$ "

## De Finetti Representation

Alternative description

1. Sample $w_{1}, \ldots, w_{K} \sim_{\mathrm{idd}} \operatorname{Beta}(1, \alpha / K)$
2. Sample $X_{1 k}, \ldots, X_{n k} \sim_{\text {iid }} \operatorname{Bernoulli}\left(w_{k}\right)$

$$
\left(\begin{array}{ccc}
w_{1} & \ldots & w_{K} \\
X_{11} & \ldots & X_{1 K} \\
\vdots & & \vdots \\
X_{N 1} & \ldots & X_{N K}
\end{array}\right)
$$

We need some form of limit object for $\operatorname{Beta}(1, \alpha / K)$ for $K \rightarrow \infty$.

## Beta Process (BP)

Distribution on objects of the form

$$
\theta=\sum_{k=1}^{\infty} w_{k} \delta_{\phi_{k}} \quad \text { with } w_{k} \in[0,1] .
$$



- IBP matrix entries are sampled as $X_{n k} \sim_{\mathrm{iid}} \operatorname{Bernoulli}\left(w_{k}\right)$.
- Beta process is the de Finetti measure of the IBP, that is, $Q=\mathrm{BP}$.
- $\theta$ is a random measure (but not normalized)


## REFERENCES I

[FLP12] S. Favaro, A. Lijoi, and I. Prünster. Conditional formulae for Gibbs-type exchangeable random partitions. Ann. Appl. Probab. To appear., 2012.
[GG06] T. L. Griffiths and Z. Ghahramani. Infinite latent feature models and the Indian buffet process. In Advances in Neural Information Processing Systems, volume 18, 2006.
[GG11] T. L. Griffiths and Z. Ghahramani. The Indian buffet process: An introduction and review. J. Mach. Learn. Res., 12:1185-1224, 2011.
[GHP07] A. V. Gnedin, B. Hansen, and J. Pitman. Notes on the occupancy problem with infinitely many boxes: General asymptotics and power laws. Probability Surveys, 4:146-171, 2007.
[GP06a] A. Gnedin and J. Pitman. Exchangeable Gibbs partitions and Stirling triangles. Journal of Mathematical Sciences, 138(3):5674-5684, 2006.
[GP06b] A. Gnedin and J. Pitman. Exchangeable Gibbs partititions and Stirling triangles. J. Math. Sci., 138(3):5674-5685, 2006.
[Hjo90] N. L. Hjort. Nonparametric Bayes estimators based on beta processes in models for life history data. Ann. Statist., 18:1259-1294, 1990.
[IJ01] H. Ishwaran and L. F. James. Gibbs sampling methods for stick-breaking priors. Journal of the American Statistical Association, 96(453): 161-173, 2001.
[JLP09] L. F. James, A. Lijoi, and I. Prüenster. Posterior analysis for normalized random measures with independent increments. Scandinavian Journal of Statistics, 36:76-97, 2009.
[Kal05] Olav Kallenberg. Probabilistic Symmetries and Invariance Principles. Springer, 2005.
[Kin75] J. F. C. Kingman. Random discrete distributions. Journal of the Royal Statistical Society, 37:1-22, 1975.
[LMP05a] A. Lijoi, R. H. Mena, and I. Prüenster. Hierarchical mixture modelling with normalized inverse-Gaussian priors. Journal of the American Statistical Association, 100:1278-1291, 2005.
[LMP05b] A. Lijoi, R. H. Mena, and I. Prünster. Hierarchical mixture modeling with normalized inverse-Gaussian priors. J. Amer: Statist. Assoc., 100:1278-1291, 2005.
[LP10] A. Lijoi and I. Prünster. Models beyond the Dirichlet process. In N. L. Hjort, C. Holmes, P. Müller, and S. G. Walker, editors, Bayesian Nonparametrics. Cambrdige University Press, 2010.
[Nea00] R. M. Neal. Markov chain sampling methods for Dirichlet process mixture models. Journal of Computational and Graphical Statistics, 9:249-265, 2000.

## REFERENCES II

[Pem07] R. Pemantle. A survey of random processes with reinforcement. Probab. Surv., 4:1-79, 2007.
[Pit03] J. Pitman. Poisson-Kingman partitions. In D. R. Goldstein, editor, Statistics and Science: a Festschrift for Terry Speed, pages 1-34. Institute of Mathematical Statistics, 2003.
[Rob95] C. P. Robert. Mixtures of distributions: inference and estimation. In W. R. Gilks, S. Richardson, and D. J. Spiegelhalter, editors, Markov Chain Monte Carlo in Practice. Chapman \& Hall, 1995.
[Sch95] M. J. Schervish. Theory of Statistics. Springer, 1995.
[Teh06] Y. W. Teh. A hierarchical Bayesian language model based on Pitman-Yor processes. In Proceedings of the 21st International Conference on Computational Linguistics and 44th Annual Meeting of the Association for Computational Linguistics, pages 985-992, 2006.
[TJ07] R. Thibaux and M. I. Jordan. Hierarchical beta processes and the Indian buffet process. In J. Mach. Learn. Res. Proceedings (AISTATS), volume 2, pages 564-571, 2007.

